

ROBUST LOCATION AND SCALE ESTIMATION WITH CENSORED OUTCOMES

Jerome H. Friedman

Stanford University

MACHINE LEARNING

$$y = F(\mathbf{x}, \mathbf{z})$$

y = outcome variable

$\mathbf{x} = (x_1, \dots, x_p)$ observed predictor variables

$\mathbf{z} = (z_1, z_2, \dots)$ other variables

Goal: estimate $E[y | \mathbf{x}]$ given data $\{y_i, \mathbf{x}_i\}_{i=1}^N$

STATISTICAL MODEL

$$y = f(\mathbf{x}) + s(\mathbf{x}) \cdot \epsilon$$

$f(\mathbf{x}) = E[y | \mathbf{x}]$ location function

$s(\mathbf{x}) > 0$ scale function

ϵ = random variable, $E[\epsilon | \mathbf{x}] = 0$

Prediction: $\hat{y} = f(\mathbf{x})$

$s(\mathbf{x}) \cdot \epsilon$ = “irreducible error” (unavoidable)

REDUCIBLE ERROR

$$r(\mathbf{x}) = E | f(\mathbf{x}) - \hat{f}(\mathbf{x}) |$$

$f(\mathbf{x})$ = optimal location (target) function

$\hat{f}(\mathbf{x})$ = estimate based on training data & ML method

ML goal: methods to reduce $r(\mathbf{x})$

Statistics goal: methods to estimate $r(\mathbf{x})$

Prediction error (y) = Reducible + Irreducible

Usually: Irreducible $s(\mathbf{x}) >>$ Reducible $r(\mathbf{x})$

USUAL ASSUMPTIONS

$s(\mathbf{x}) = s = \text{constant}$ (homoscedasticity)

$\epsilon \sim N(0, 1)$ (normality)

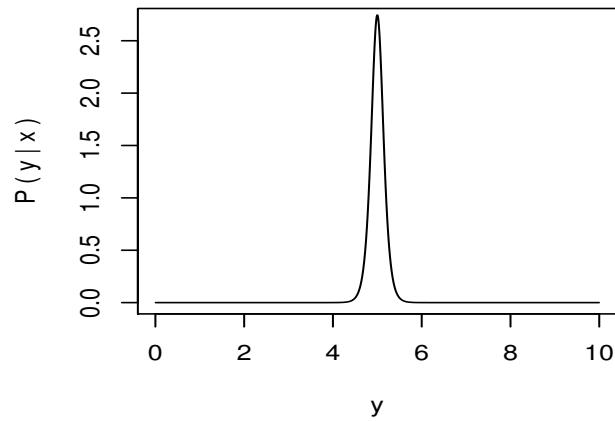
HOMOSCEDASTICITY

$$F(\mathbf{x}, \mathbf{z}) = f(\mathbf{x}) + g(\mathbf{z}) \text{ additive}$$

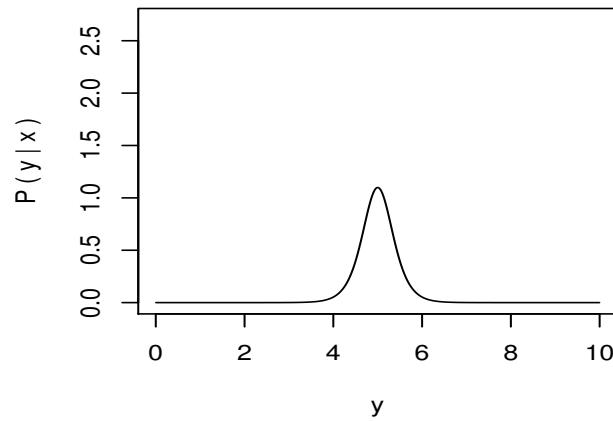
$$p(\mathbf{x}, \mathbf{z}) \implies \text{scale } [g(\mathbf{z}) \mid \mathbf{x}] = \text{constant}$$

Not very likely

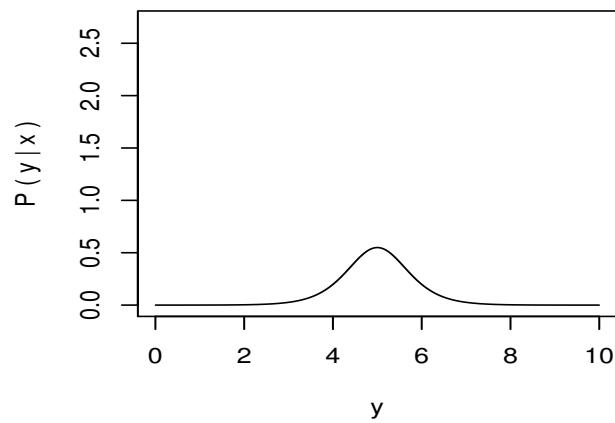
loc = 5, scale = 0.1



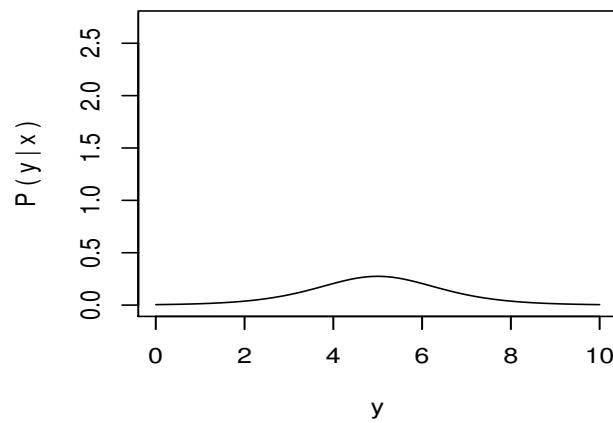
loc = 5, scale = 0.25



loc = 5, scale = 0.5



loc = 5, scale = 1



NORMALITY - not very likely either

Tukey:

“small residuals \simeq normal, larger have heavier tails.”

Heterodistributionality

Heterodistributionality

Robustness:

Choose compromise $\bar{p}(\epsilon)$

good properties for others

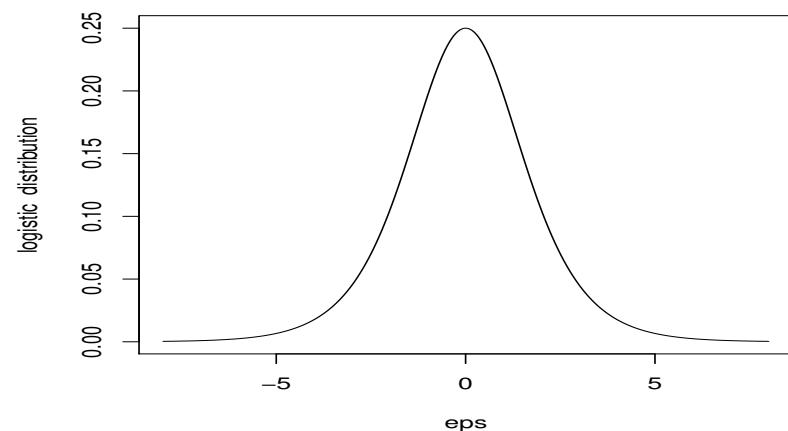
$\bar{p}(\epsilon)$ = normal, not good!

LOGISTIC DISTRIBUTION

$$\epsilon | \mathbf{x} = (y - f(\mathbf{x}))/s(\mathbf{x})$$

$$\bar{p}(\epsilon) = \frac{e^{-\epsilon}}{s(1+e^{-\epsilon})^2}$$

small $|\epsilon| \sim \text{normal}$, large $|\epsilon| \sim \text{exponential}$



Prediction: $\hat{y} = \hat{f}(\mathbf{x})$

$$\hat{f}(\mathbf{x}) = \arg \min_{f \in F} \sum_{i=1}^N [\varepsilon_i + 2 \log(1 + e^{-\varepsilon_i})]$$

$$\varepsilon_i = (y_i - f(\mathbf{x}_i))/s(\mathbf{x}_i)$$

minimized at $f(\mathbf{x}_i) = y_i$ indep $s(\mathbf{x}_i)$

$1/s(\mathbf{x}_i)$ ~ “weight” for obs i

controls relative influence of i to fit

Using incorrect $s(x)$ to estimate $f(x)$

increases variance, not bias

assume $s(x) = \text{constant}$ usually not too bad

ESTIMATE $\hat{s}(\mathbf{x})$

(1) Improve $\hat{f}(\mathbf{x})$ in high variance settings.

(2) Important inferential statistic:

(a) prediction interval \sim accuracy of \hat{y} -prediction:

$$\text{logistic: } IQR[y \mid f(\mathbf{x})] = 2 s(\mathbf{x}) / \log(3)$$

(b) can affect decision

(3) Crucial with censoring

CENSORING (y -value partially known)

Data: $\{y_i, \mathbf{x}_i\}_1^N \rightarrow \{a_i, b_i, \mathbf{x}_i\}_1^N$

$$a_i \leq y_i \leq b_i$$

$a_i = b_i = y_i \Rightarrow y$ -value known

$a_i = -\infty \Rightarrow$ censored below b_i

$b_i = \infty \Rightarrow$ censored above a_i

Otherwise: interval censored $[a_i, b_i]$

Special Case

$\{a_i, b_i\} \rightarrow K$ disjoint intervals (bins):

$K = 2 \Rightarrow$ usual binary logistic regression

$K > 2 \Rightarrow$ ordered multiclass logistic regression

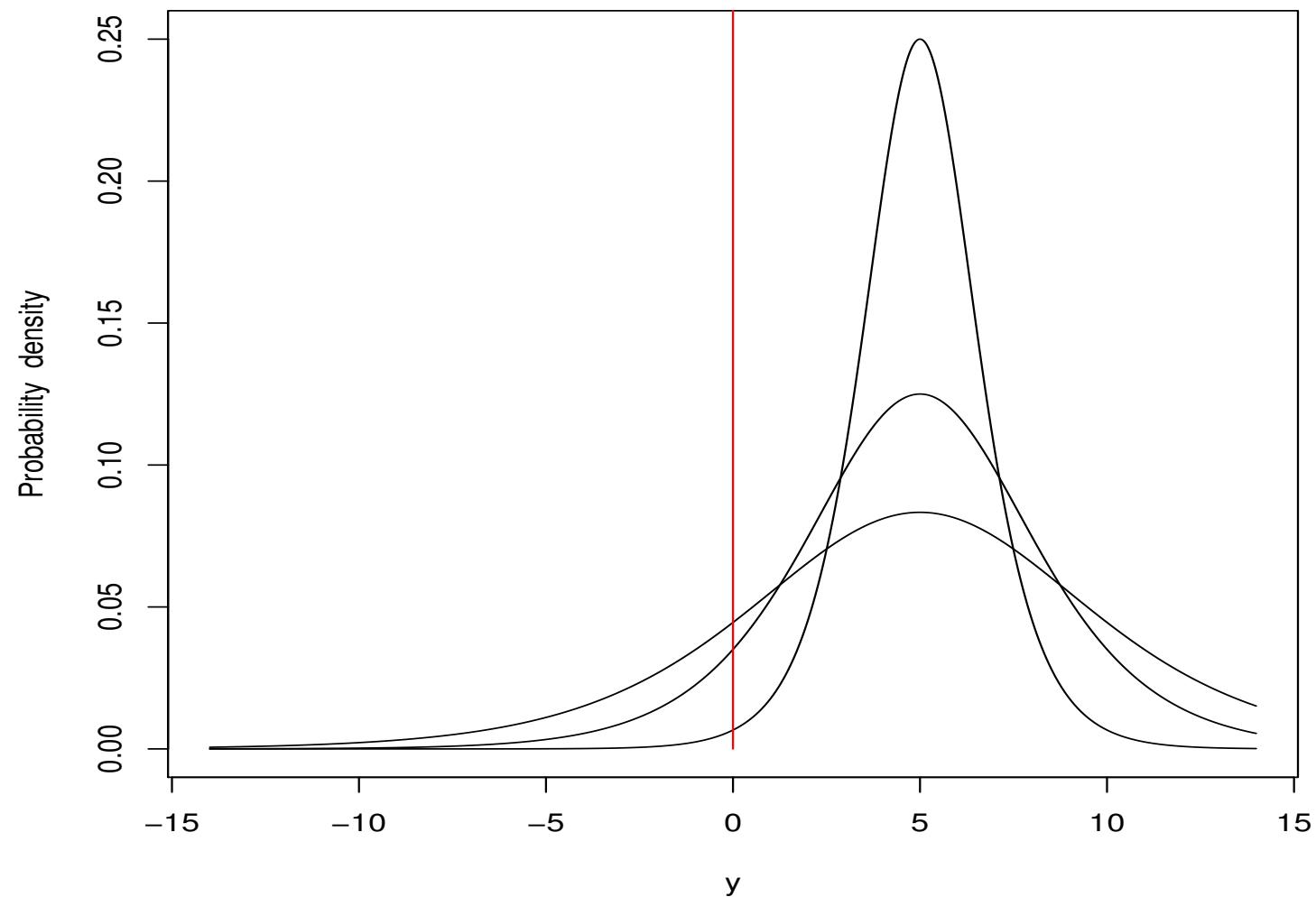
LIKELIHOOD

$$\Pr(a \leq y \leq b) = \frac{1}{1+e^{-(b-f)/s}} - \frac{1}{1+e^{-(a-f)/s}}$$

Depends strongly on *both f and s*

Need to estimate *both f(x) and s(x)*

Logistic distribution: $f = 5$



EXERCISE

$$[\hat{f}(\mathbf{x}), \hat{s}(\mathbf{x})] = \arg \min_{(f,s) \in F} \sum_{i=1}^N L[a_i, b_i, f(\mathbf{x}_i), s(\mathbf{x}_i)]$$

$$L(a, b, f, s) = -\log \left(\frac{1}{1+e^{-(b-f)/s}} - \frac{1}{1+e^{-(a-f)/s}} \right)$$

PROBLEM

$L(a, b, f, s)$ NOT convex in s

IS convex in $t = 1/s \Rightarrow$ solve for t

Constraint $t > 0 \Rightarrow$ solve for $\log(t) = -\log(s)$

GRADIENT BOOSTED TREE ENSEMBLES

Ann. Statist, **29**. 1189 – 1232 (2001)

$$\hat{f}(\mathbf{x}) = \sum_{k=1}^{K_f} T_k^{(f)}(\mathbf{x})$$

$$\log(\hat{s}(\mathbf{x})) = \sum_{k=1}^{K_s} T_k^{(s)}(\mathbf{x})$$

$$T_k(\mathbf{x}) = \text{CART-tree}(\mathbf{x})$$

ITERATIVE GRADIENT BOOSTING

Start: $\hat{s}(\mathbf{x}) = \text{constant}$

Loop {

$\hat{f}(\mathbf{x}) = \text{tree-boost } f(\mathbf{x}) \text{ given } \hat{s}(\mathbf{x})$

$\log(\hat{s}(\mathbf{x})) = \text{tree-boost } \log(s(\mathbf{x})) \text{ given } \hat{f}(\mathbf{x})$

}

Until no change

DIAGNOSTICS

$$(1) \ median [y | f(\mathbf{x})] = f(\mathbf{x})$$

$$(2) \ median [|y - f(\mathbf{x})| | s(\mathbf{x})] = s(\mathbf{x}) \cdot \log(3)$$

$$(3) \# (y_i \in [u, v] | f_i \in [g, h]) =$$

$$\sum_{f_i \in [g, h]} \left(\frac{1}{1+e^{-(v-f_i)/s_i}} - \frac{1}{1+e^{-(u-f_i)/s_i}} \right)$$

$$(f_i = \hat{f}(\mathbf{x}_i), \quad s_i = \hat{s}(\mathbf{x}_i))$$

California Housing Price Data (STATLIB Repository)

$N = 20460$ CA neighborhoods (1990 census block groups)

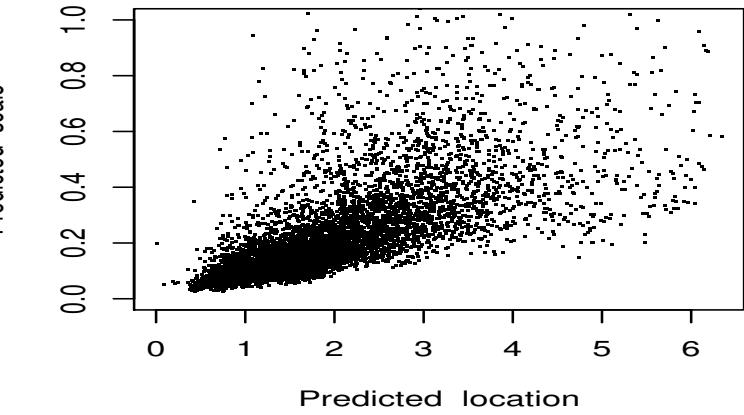
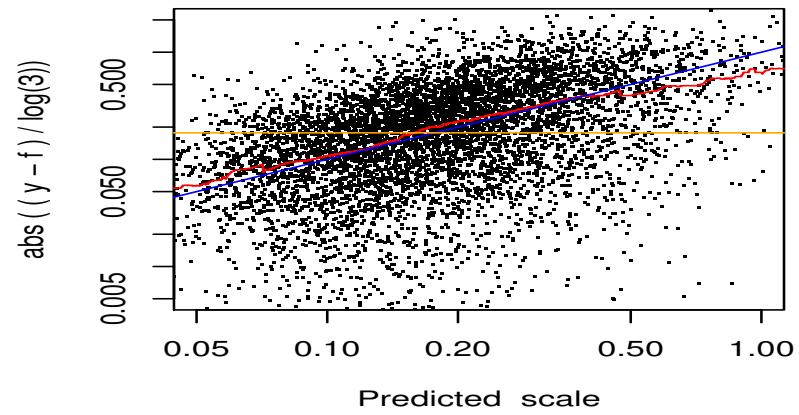
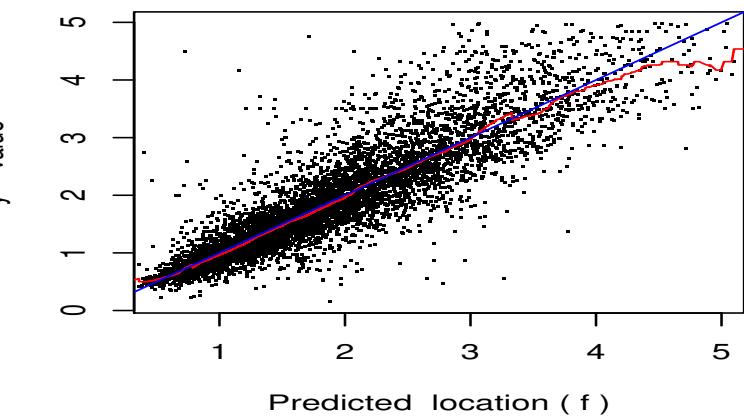
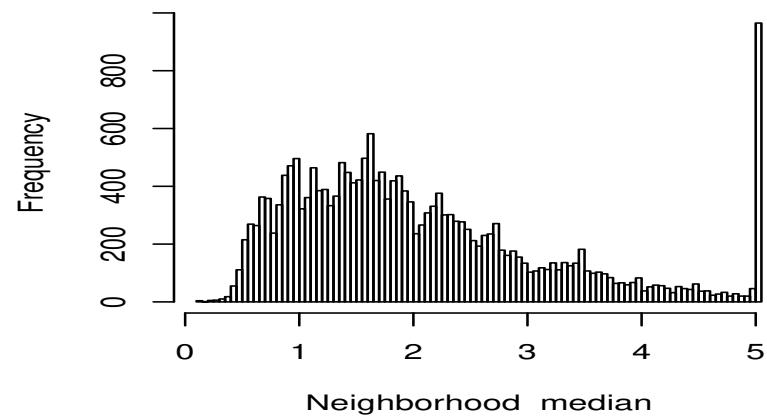
y = Median House Value

\mathbf{x} = (Median Income, Housing Median Age,

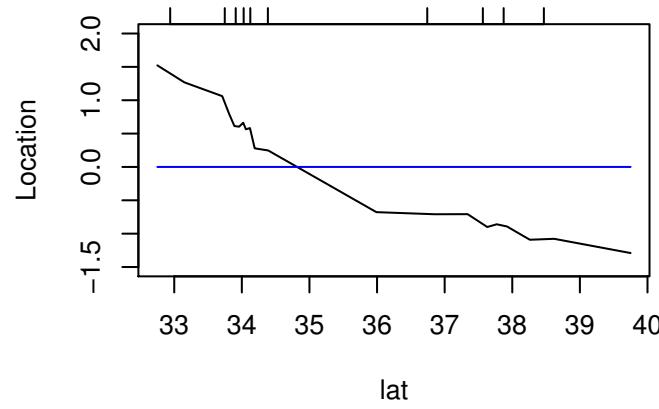
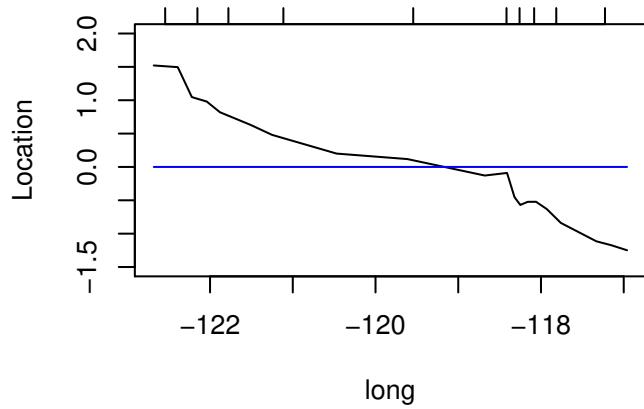
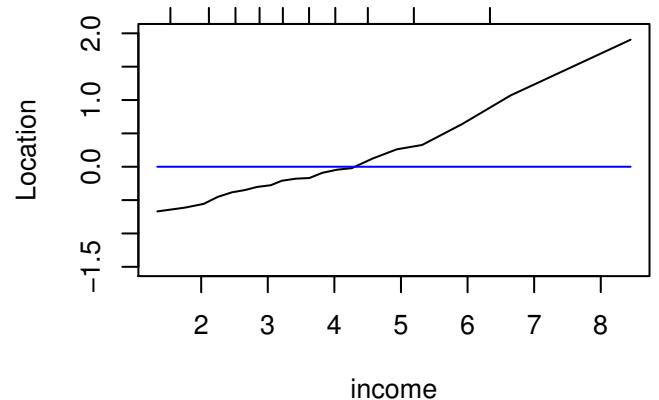
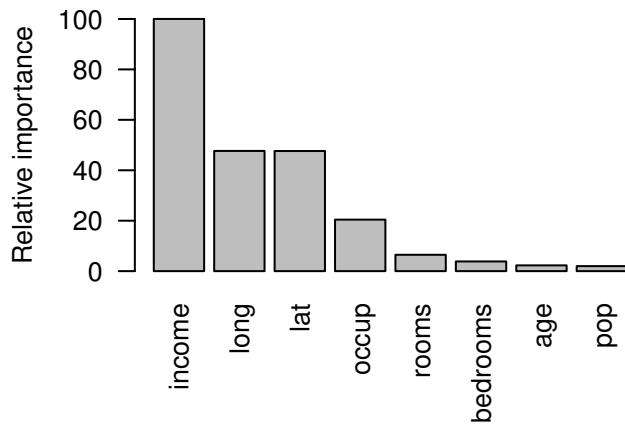
Ave No Rooms, Ave No Bedrooms,

Population, Ave Occupancy, Latitude, Longitude)

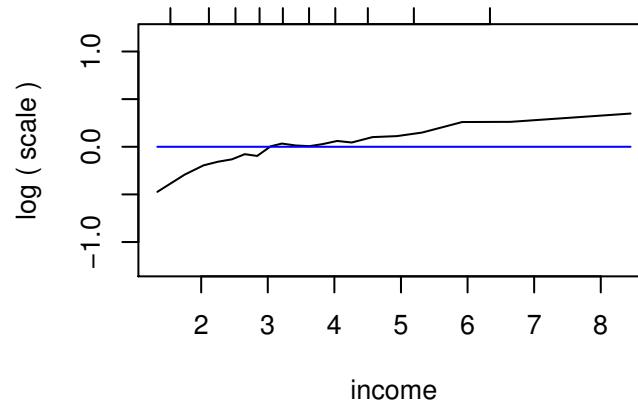
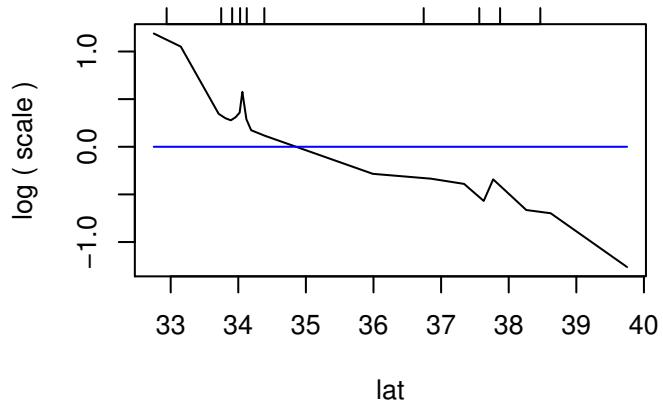
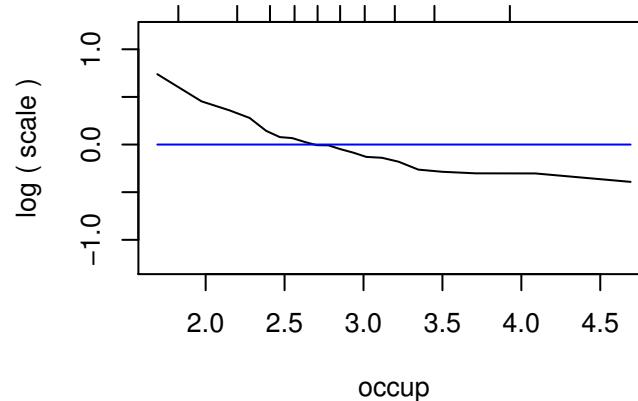
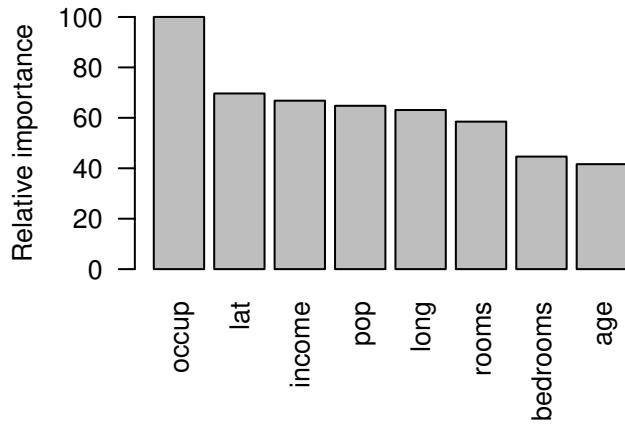
CA housing prices



CA housing : location model



CA housing : log (scale) model

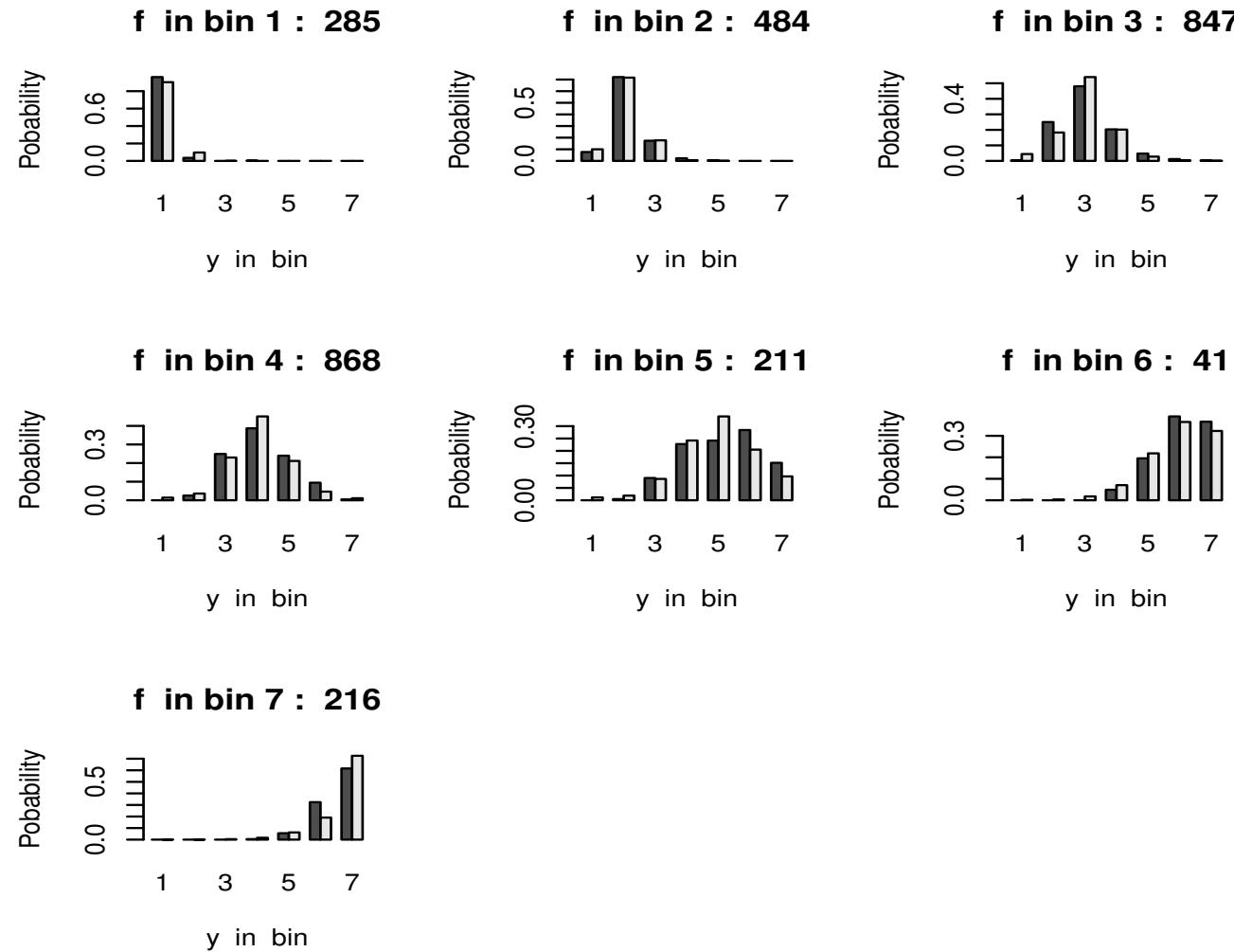


QUESTIONNAIRE DATA

$$N = 8857, \ p = 13$$

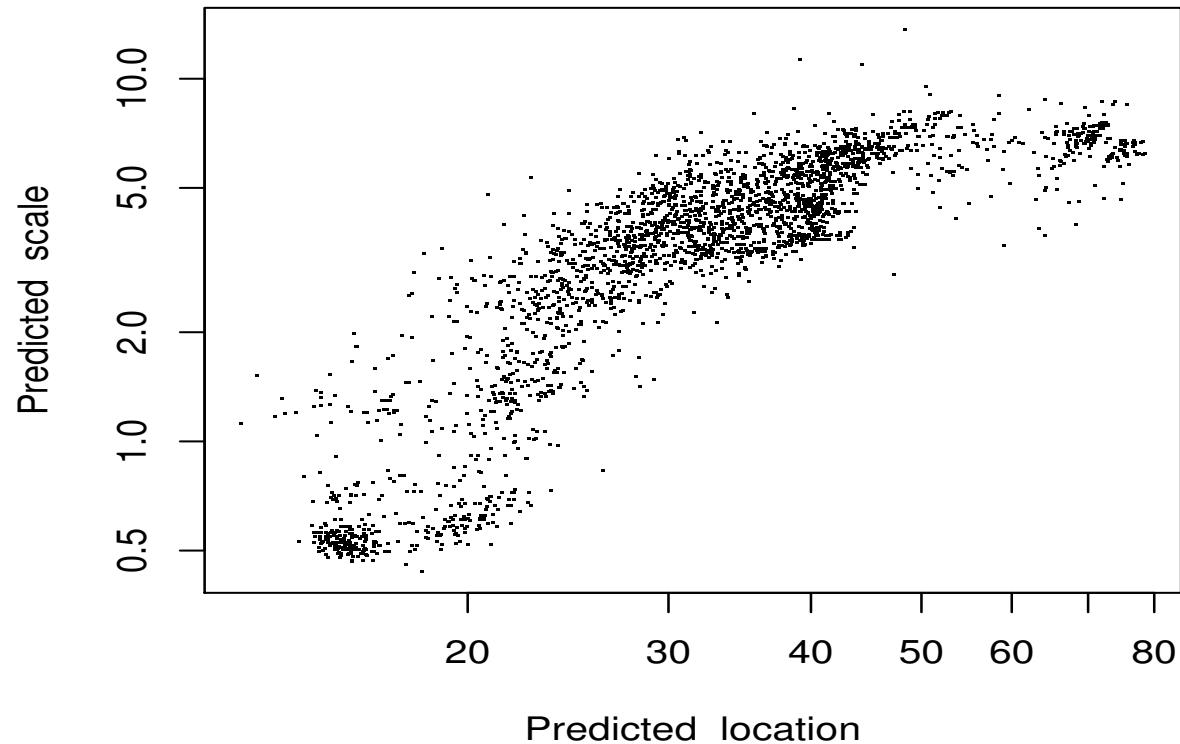
$$y = AGE \in \left\{ \begin{array}{ll} 14 & 17 \\ 18 & 24 \\ 25 & 34 \\ 35 & 44 \\ 45 & 54 \\ 55 & 64 \\ 65 & \infty \end{array} \right|$$

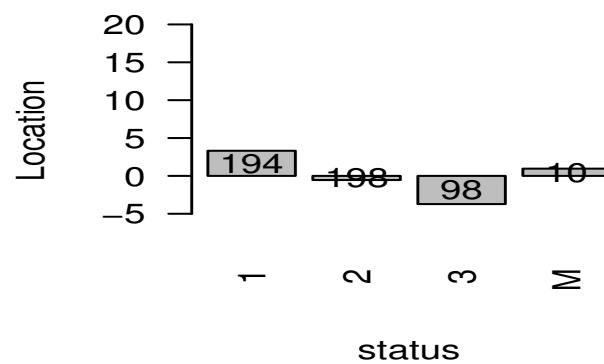
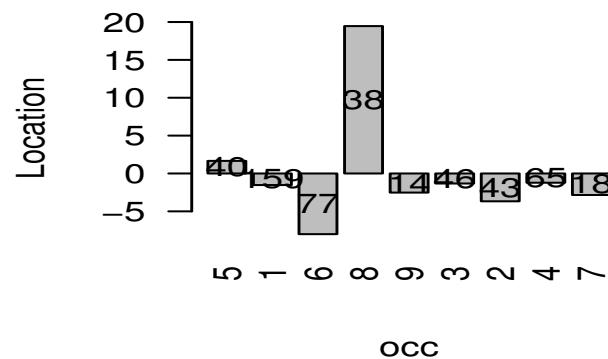
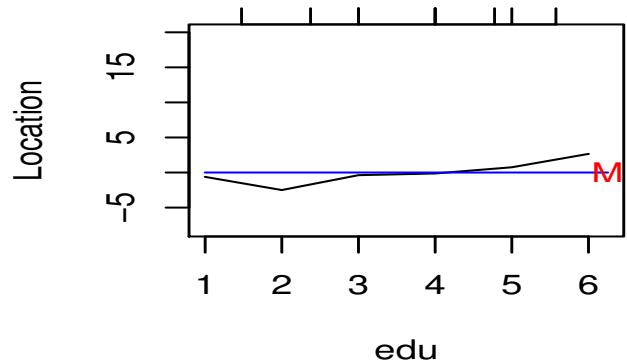
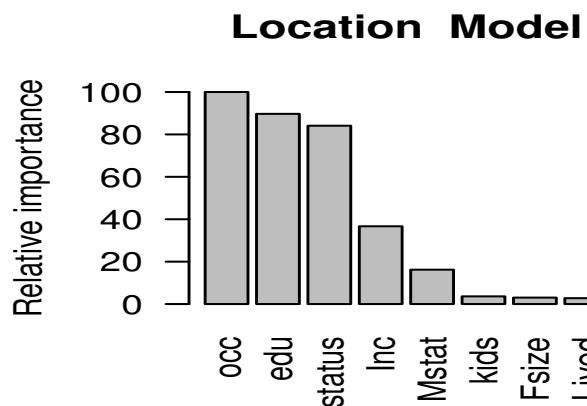
$x = (\text{Occupation}, \text{Type of Home}, \text{Sex},$
 $\text{Marital Status}, \text{Education}, \text{Income},$
 $\text{Lived in BA}, \text{Dual Incomes}, \text{Persons in}$
 $\text{Household}, \text{Persons in Household} < 18,$
 $\text{Householder Status}, \text{Ethnicity}, \text{Language})$

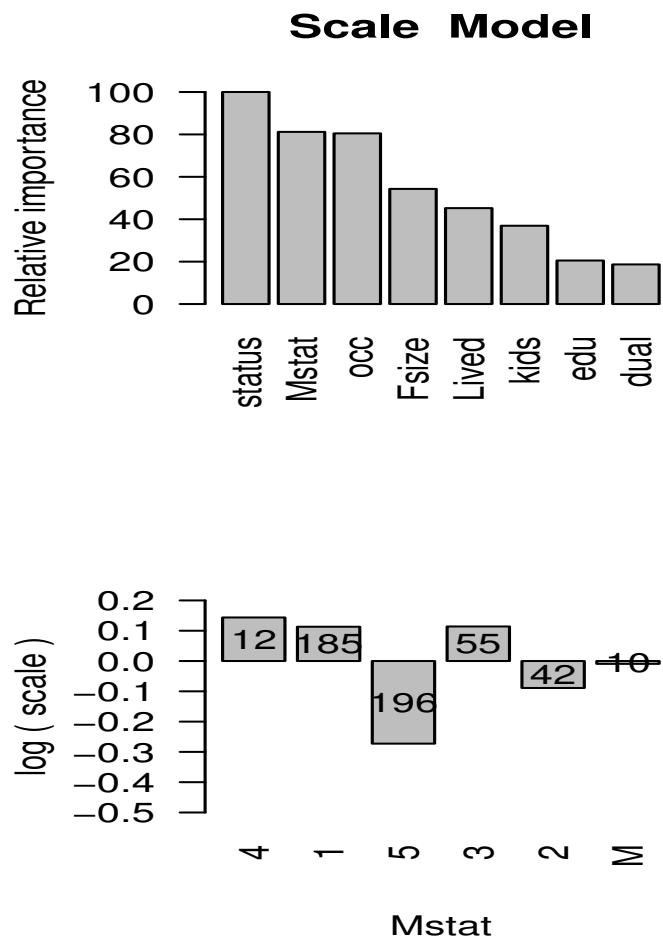


AGE predictions

AGE







Wine Quality Data (Irvine Repository)

$N = 6497$ samples of Portuguese "Vinho Verde"

$\tilde{y} = \text{Quality: integer } (1, 2, \dots, 10)$

median of at least 3 expert evaluations

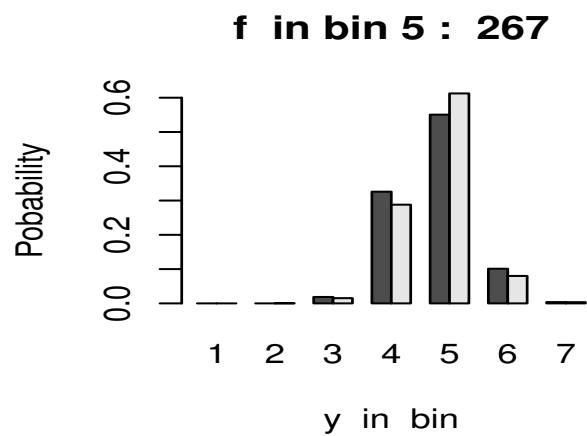
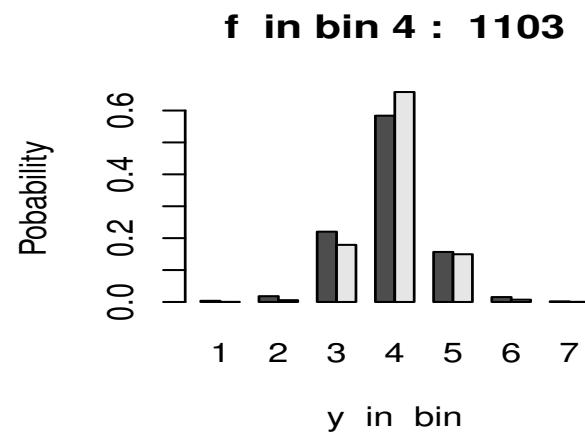
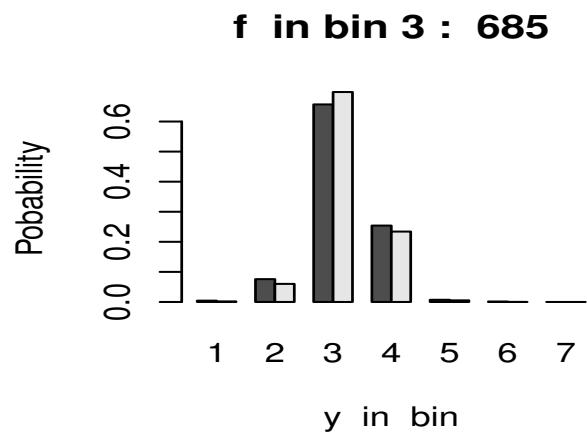
$\tilde{y} = k \Rightarrow y \in [k - 1/2, k + 1/2]$

$x =$ (Fixed acidity, Volatile acidity, Citric acid,

Residual sugar, Chlorides, Free sulfur dioxide

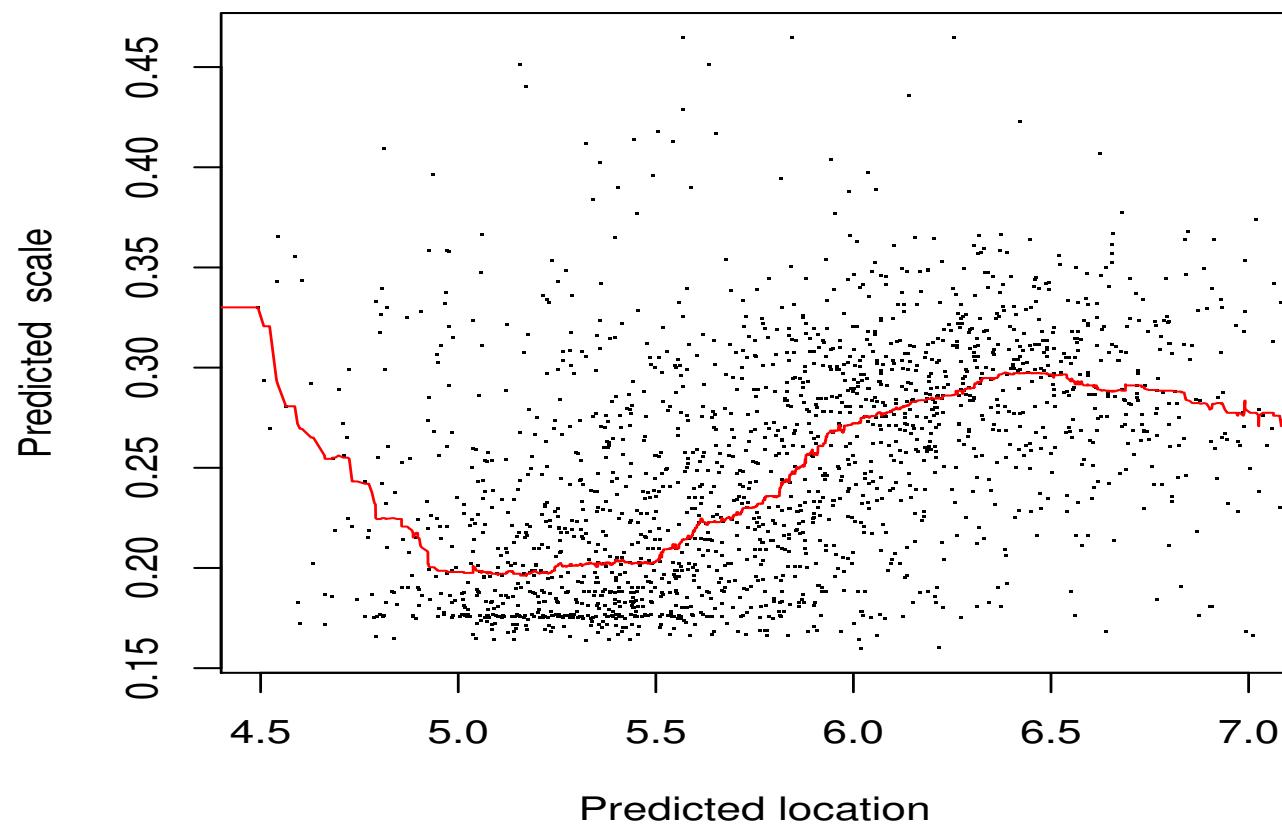
Total sulfur dioxide, Density, pH, Sulfates,

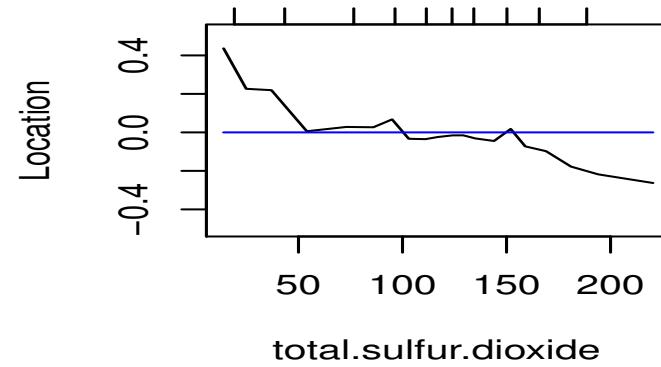
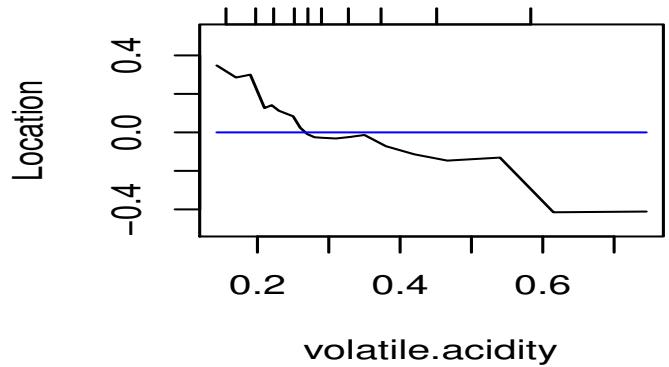
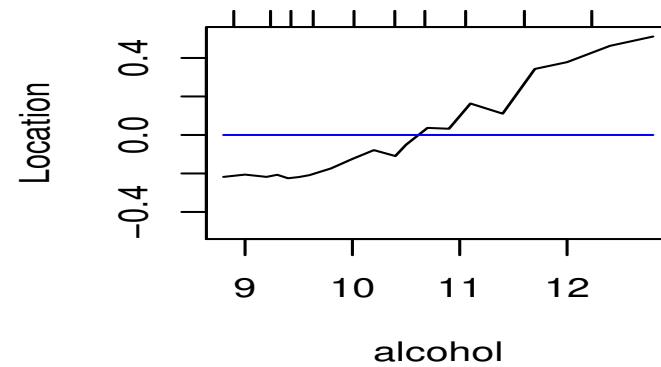
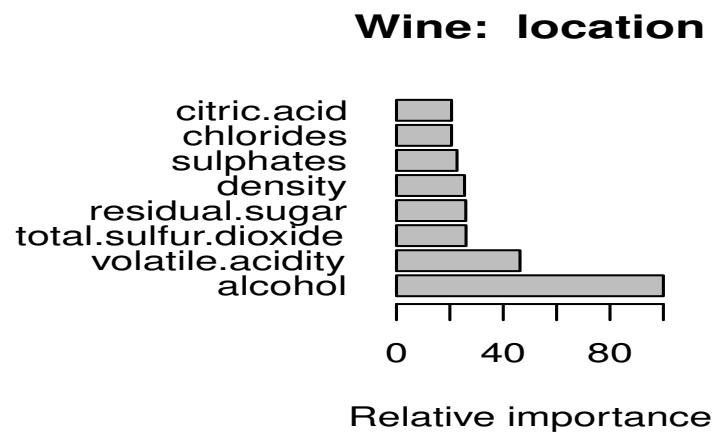
Alcohol)

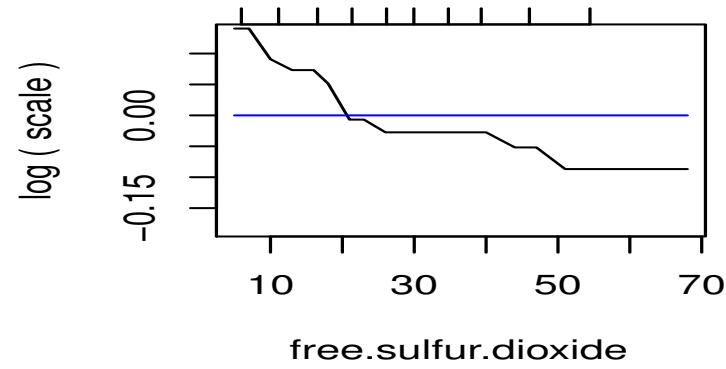
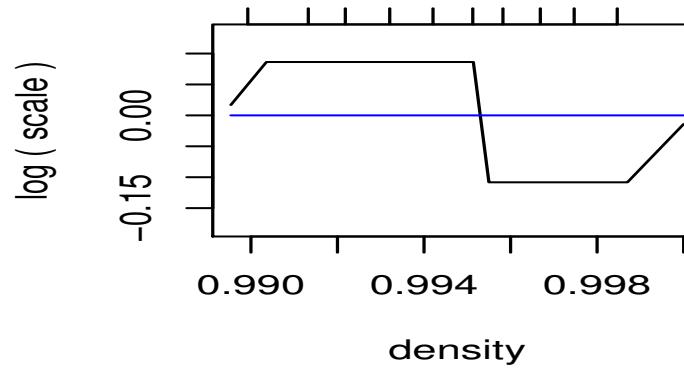
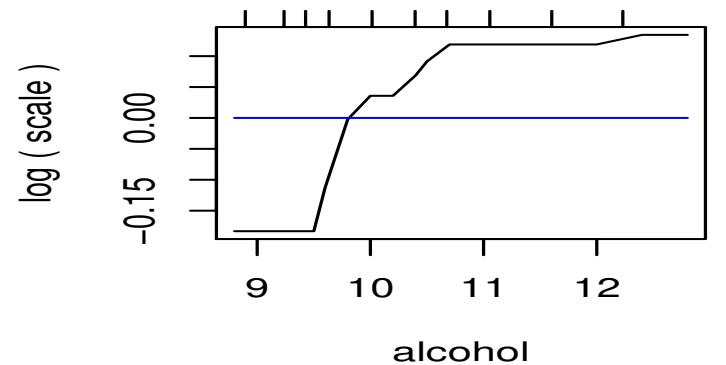
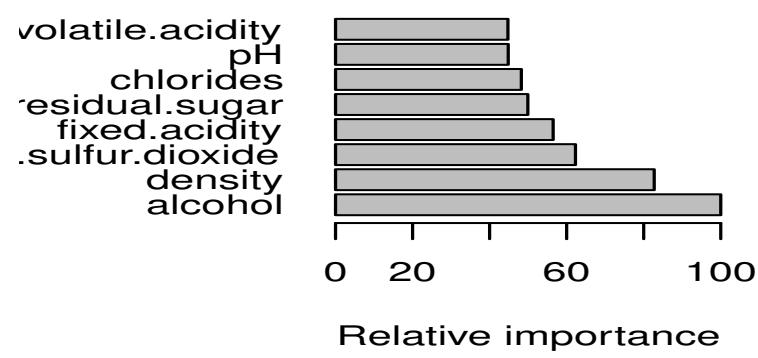


Wine quality data

Wine quality data







ORDERED MULTICLASS LOGISTIC REGRESSION

$$y_i \in \{C_1 < C_2, \dots, C_{K-1} < C_K\}$$

Interval censored:

$\{a_i, b_i\} \rightarrow K$ disjoint intervals (bins):

$$\{b_0, b_1, \dots, b_K\} \quad b_0 = -\infty, \quad b_K = \infty$$

bins \sim classes with separating boundaries

$$\mathbf{b} = \{b_1, b_2, \dots, b_{K-1}\} \quad \text{unknown}$$

(overall location & scale arbitrary)

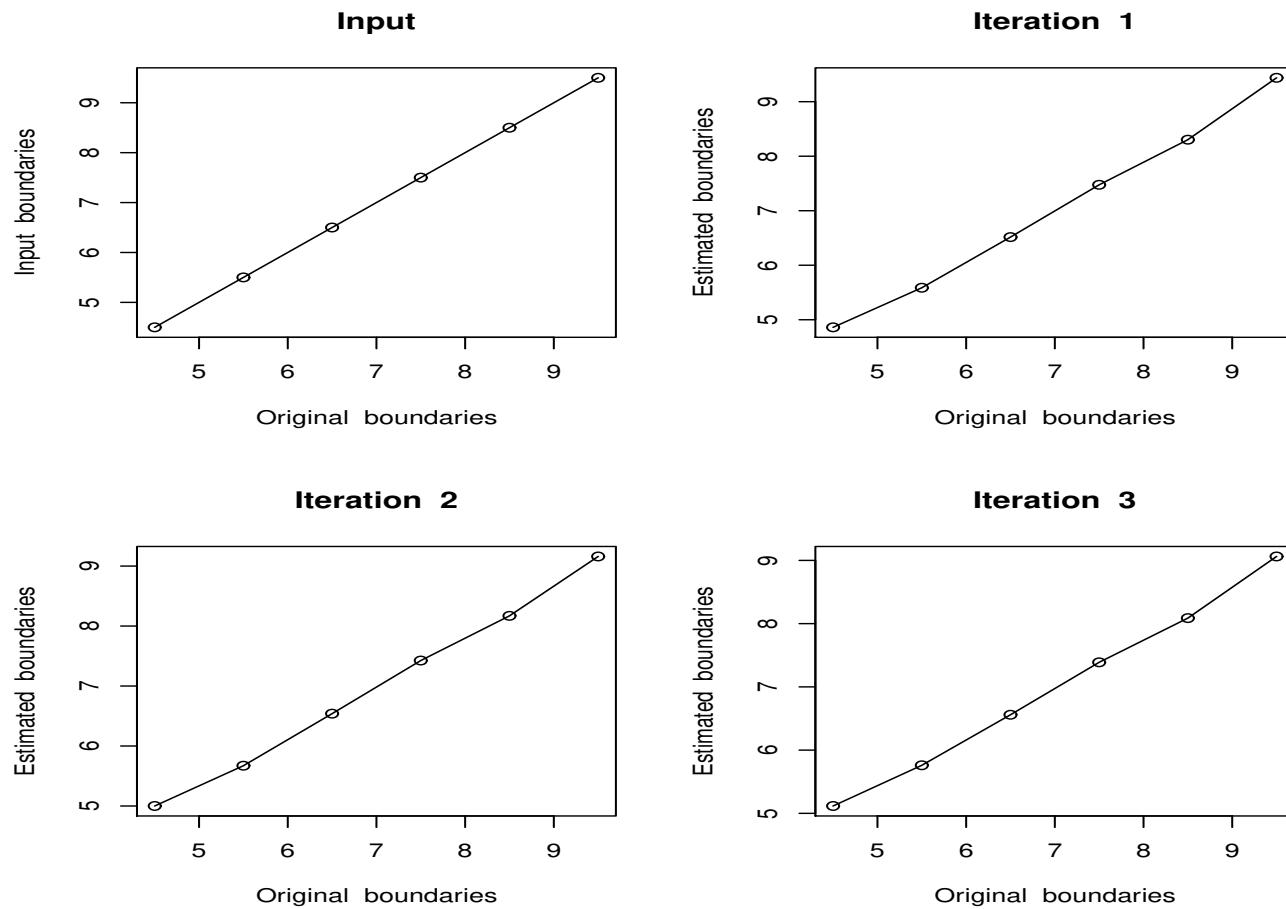
OPTIMAL SCALING (aka ACE)

$$[\hat{\mathbf{b}}, \hat{f}(\mathbf{x}), \hat{s}(\mathbf{x})] = \arg \min_{\mathbf{b}, (f, s) \in F} \sum_{i=1}^N L[b_{k(i)-1}, b_{k(i)}, f(\mathbf{x}_i), s(\mathbf{x}_i)]$$

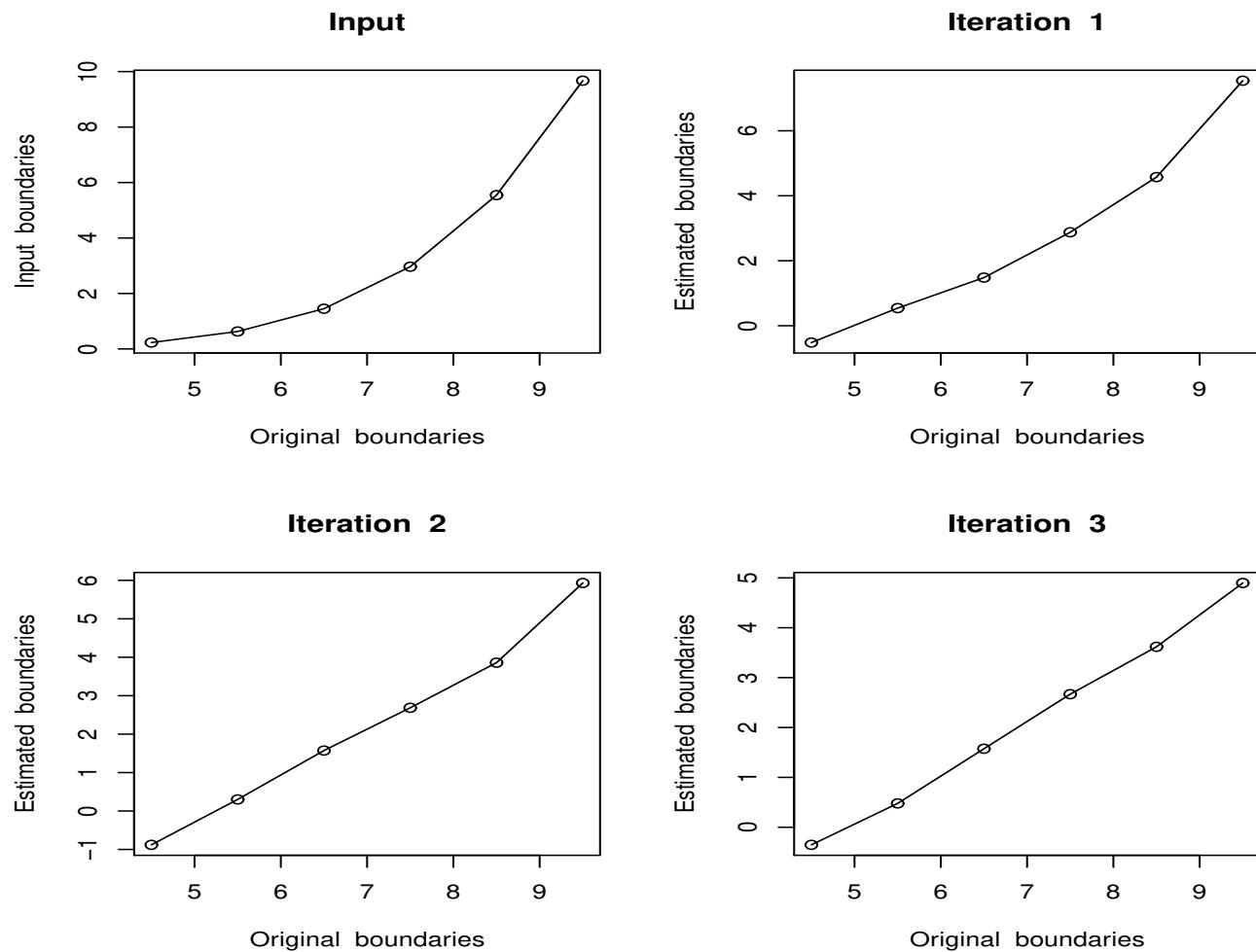
$$L(u, v, f, s) = -\log \left(\frac{1}{1+e^{(f-v)/s}} - \frac{1}{1+e^{(f-u)/s}} \right)$$

Alternating optimization:

$$\mathbf{b} : \sum_{i=1}^N \left(\frac{1}{1 + e^{(f(\mathbf{x}_i) - b_k)/s(\mathbf{x}_i)}} - \sum_{j=1}^k I(y_i = c_j) \right) = 0$$



Wine quality data – optimal scaling



Wine quality data – optimal scaling

ASYMMETRIC ERRORS

$$y | \mathbf{x} = f(\mathbf{x}) + \begin{cases} s_l(\mathbf{x}) \cdot \varepsilon & \varepsilon \leq 0 \\ s_u(\mathbf{x}) \cdot \varepsilon & \varepsilon > 0 \end{cases}$$

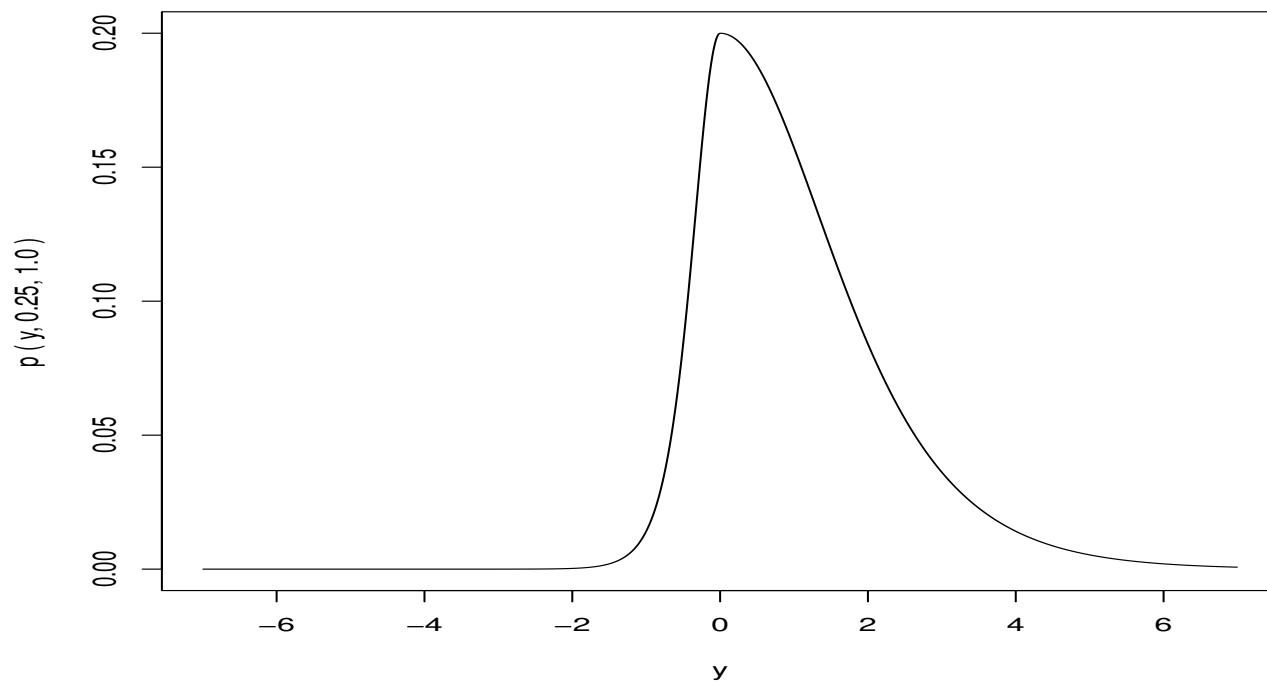
$f(\mathbf{x})$ = mode at \mathbf{x}

$s_l(\mathbf{x})$ = lower scale at \mathbf{x}

$s_u(\mathbf{x})$ = upper scale at \mathbf{x}

ε ~ standard logistic distribution

Asymmetric logistic



$$f = 0, \quad s_l = 1/4, \quad s_u = 1$$

EXERCISE (no censoring)

$$[\hat{f}(\mathbf{x}), \hat{s}_l(\mathbf{x}), \hat{s}_u(\mathbf{x})]$$

$$= \arg \min_{(f, s_l, s_u) \in F} \sum_{i=1}^N L[y_i, f(\mathbf{x}_i), s_l(\mathbf{x}_i), s_u(\mathbf{x}_i)]$$

$$L[y, f, s_l, s_u] =$$

$$L[y, f, s_l] \cdot I[y - f \leq 0] + L[y, f, s_u] \cdot I[y - f > 0]$$

$$L(y, f, s) = \log(s) + (y - f)/s + 2 \log(1 + e^{-(y-f)/s})$$

Iterative gradient boosting

ASYMMETRIC DIAGNOSTICS

$$(1) \ median [y | f(\mathbf{x}), s_l(\mathbf{x}), s_u(\mathbf{x})] = f(\mathbf{x})$$

$$+ \begin{cases} s_u(\mathbf{x}) \log \left(\frac{3s_u(\mathbf{x}) - s_l(\mathbf{x})}{s_u(\mathbf{x}) + s_l(\mathbf{x})} \right) & s_l(\mathbf{x}) \leq s_u(\mathbf{x}) \\ -s_l(\mathbf{x}) \log \left(\frac{3s_l(\mathbf{x}) - s_u(\mathbf{x})}{s_u(\mathbf{x}) + s_l(\mathbf{x})} \right) & s_l(\mathbf{x}) > s_u(\mathbf{x}) \end{cases}$$

$$(2) \ median_{y \leq f(\mathbf{x})} [|y - f(\mathbf{x})| | s_l(\mathbf{x})] / \log(3) = s_l(\mathbf{x})$$

$$(3) \ median_{y > f(\mathbf{x})} [|y - f(\mathbf{x})| | s_u(\mathbf{x})] / \log(3) = s_u(\mathbf{x})$$

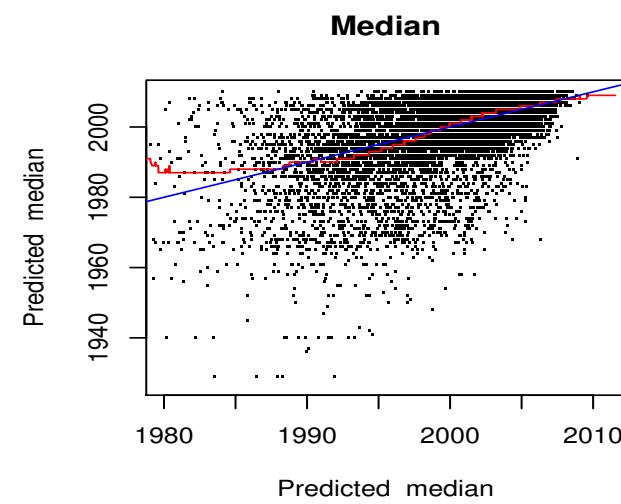
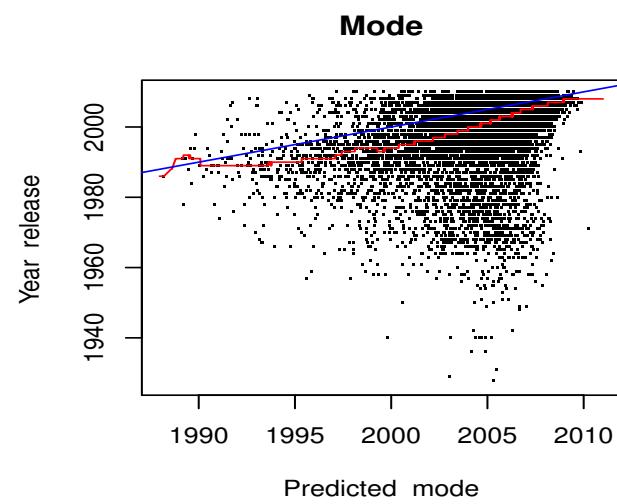
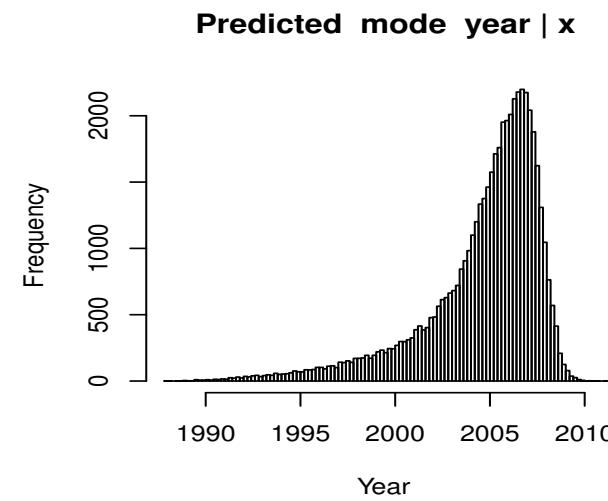
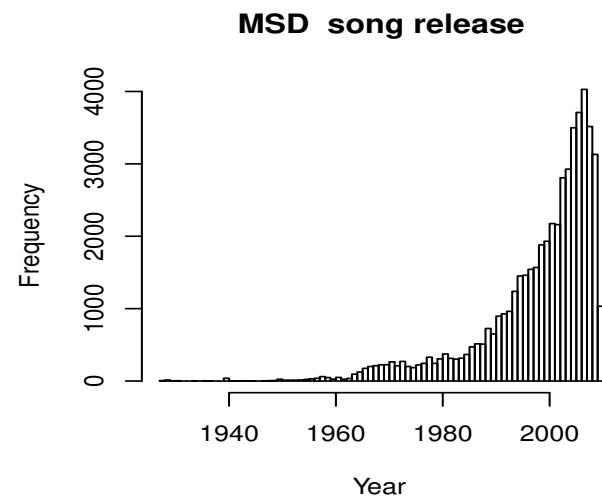
(1/2) – Million Song Dataset (Irvine Repository)

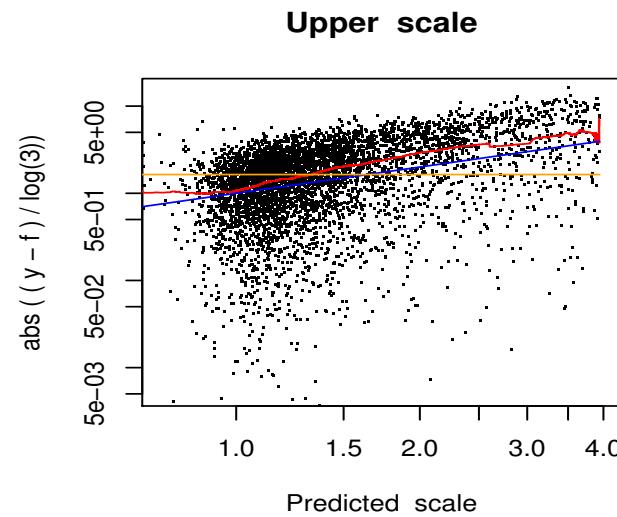
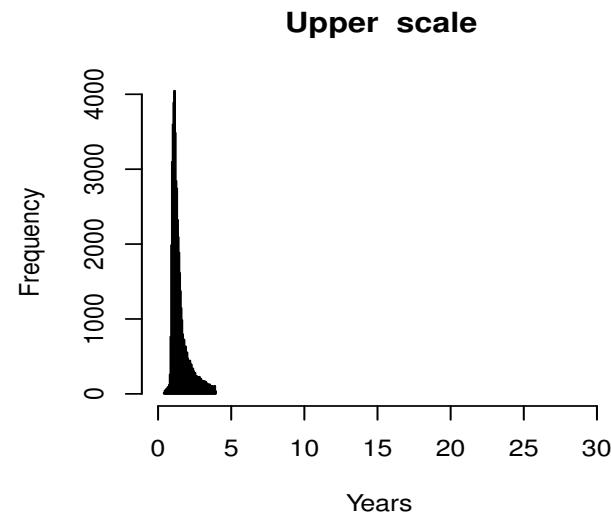
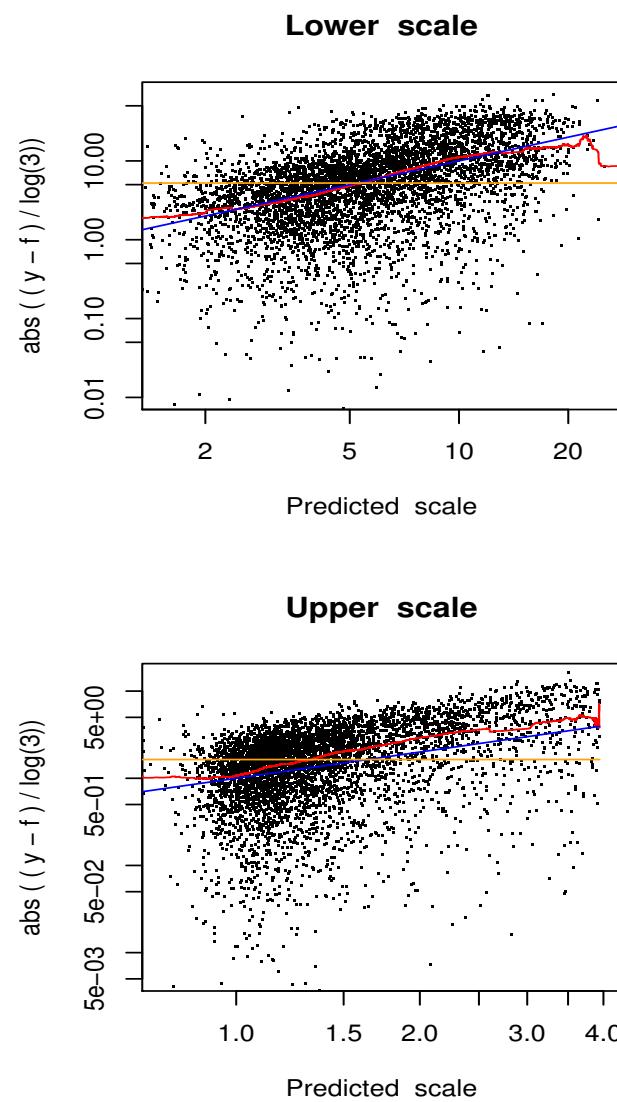
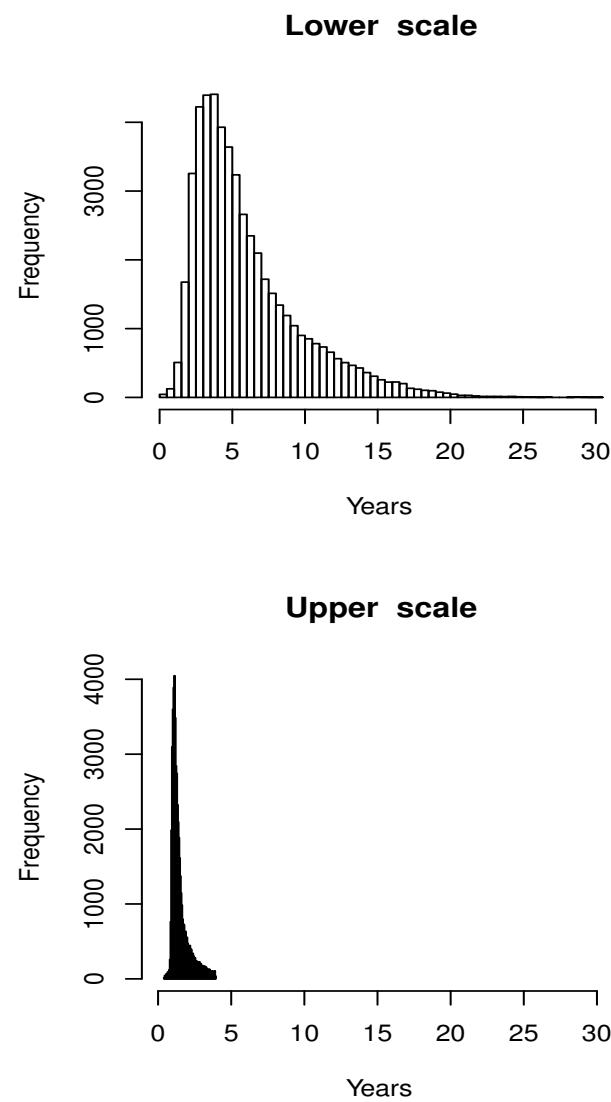
$N = 515345$ songs (463715 train, 51630 test)

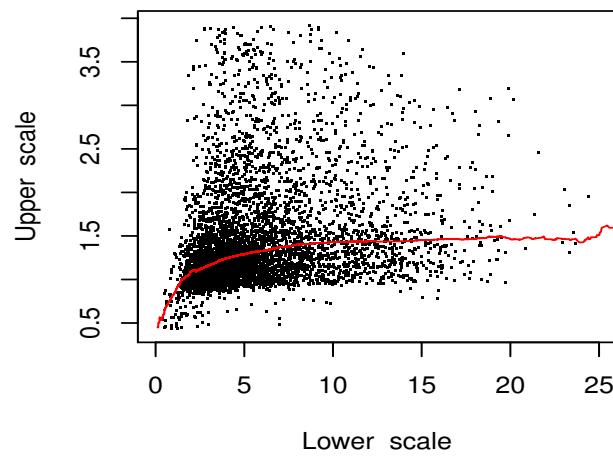
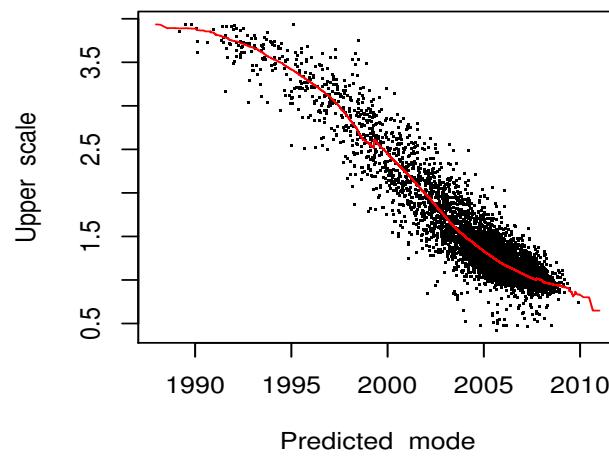
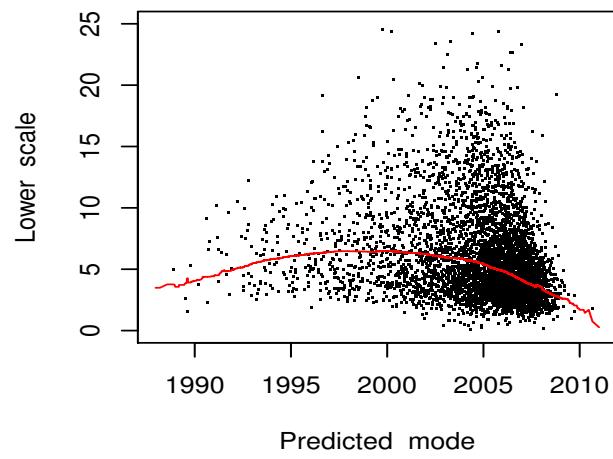
$y =$ year released (1922 – 2011)

$\mathbf{x} = 90$ attributes (Echo Nest API):

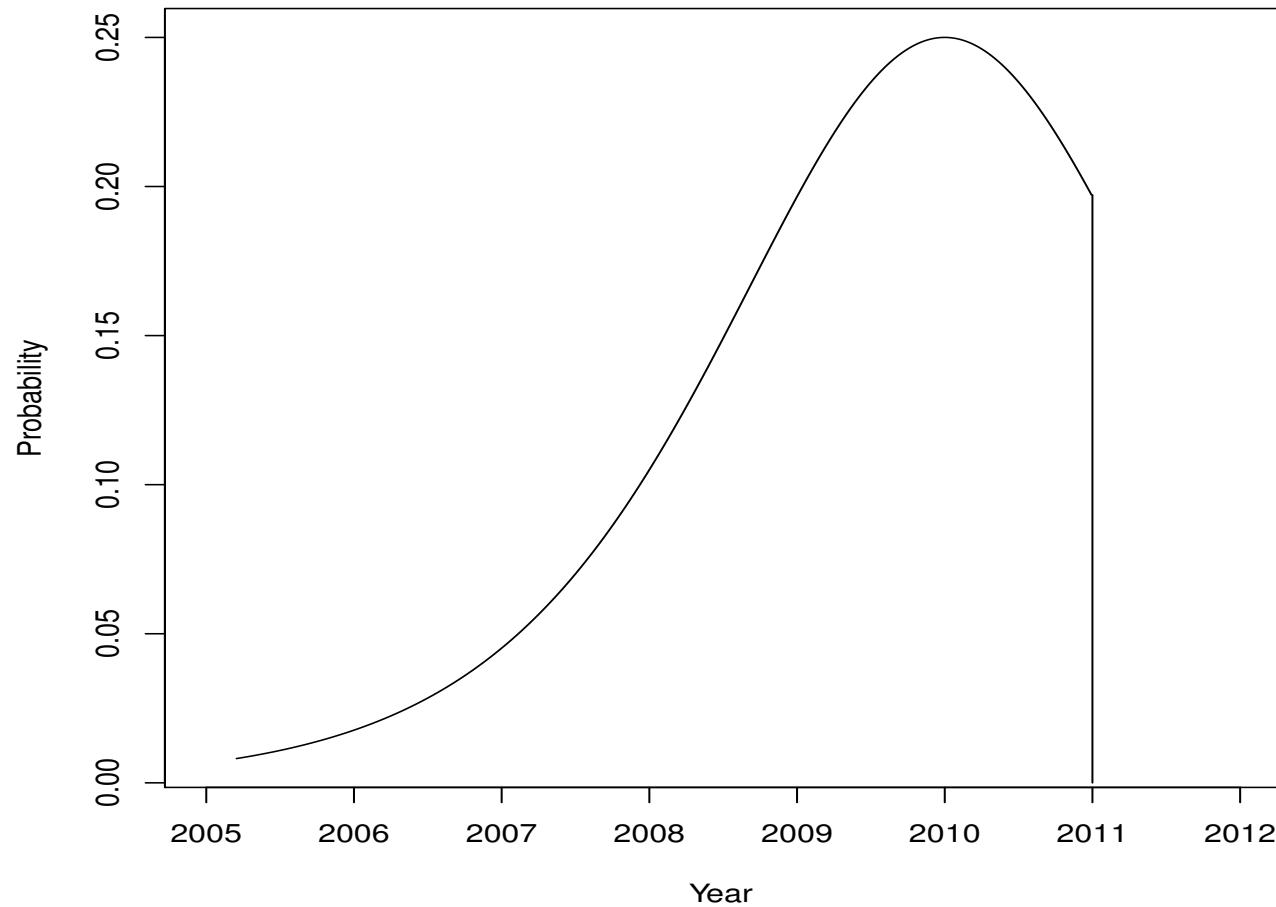
12 timbre average, 78 = timbre covariance







Truncated logistic



Predict $y \mid \mathbf{x}$, settle for $p(y \mid \mathbf{x})$

1. ROBUST (logistic) estimate: loc $f(\mathbf{x})$ & scale $s(\mathbf{x})$
2. General censoring: $a_i \leq y_i \leq b_i$ (ubiquitous)
3. Graphical diagnostics
4. Ordered multi-class classification
5. Asymmetric $p(y \mid \mathbf{x})$: $f(\mathbf{x})$, $s_l(\mathbf{x})$, $s_u(\mathbf{x})$

REFERENCES

Gradient boosting: Ann. Statist, **29**. 1189 – 1232 (2001)

Optimal scaling:

ACE: J. Amer. Statist. Assoc. **80**. 580 – 598 (1985)

GIFI: *Nonlinear Multivariate Analysis*. Wiley, N.Y. (1990)

Slides: <http://statweb.stanford.edu/~jhf/talks/kdd.pdf>